



Seasonality Shows Evidence for Polygenic Architecture and Genetic Correlation With Schizophrenia and Bipolar Disorder

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ABSTRACT

Objective: To test common genetic variants for association with seasonality (seasonal changes in mood and behavior) and to investigate whether there are shared genetic risk factors between psychiatric disorders and seasonality.

Method: Genome-wide association studies (GWASs) were conducted in Australian (between 1988 and 1990 and between 2010 and 2013) and Amish (between May 2010 and December 2011) samples in whom the Seasonal Pattern Assessment Questionnaire (SPAQ) had been administered, and the results were meta-analyzed in a total sample of 4,156 individuals. Genetic risk scores based on results from prior large GWAS studies of bipolar disorder, major depressive disorder (MDD), and schizophrenia were calculated to test for overlap in risk between psychiatric disorders and seasonality.

Results: The most significant association was with rs11825064 ($P = 1.7 \times 10^{-6}$, $\beta = 0.64$, standard error = 0.13), an intergenic single nucleotide polymorphism (SNP) found on chromosome 11. The evidence for overlap in risk factors was strongest for schizophrenia and seasonality, with the schizophrenia genetic profile scores explaining 3% of the variance in log-transformed global seasonality scores. Bipolar disorder genetic profile scores were also associated with seasonality, although at much weaker levels (minimum P value = 3.4×10^{-3}), and no evidence for overlap in risk was detected between MDD and seasonality.

Conclusions: Common SNPs of large effect most likely do not exist for seasonality in the populations examined. As expected, there were overlapping genetic risk factors for bipolar disorder (but not MDD) with seasonality. Unexpectedly, the risk for schizophrenia and seasonality had the largest overlap, an unprecedented finding that requires replication in other populations and has potential clinical implications considering overlapping cognitive deficits in seasonal affective disorders and schizophrenia.

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Although seasonal changes in mood and behavior (seasonality) have been recognized for a long time,¹ seasonal affective disorder (SAD) was first defined by Rosenthal et al² as a syndrome with recurrent depression in fall and winter and alleviation of depressive symptoms in spring and summer. In addition, while SAD is not a distinct clinical entity in the fourth and fifth iterations of the *Diagnostic and Statistical Manual of Mental Disorders (DSM)*, these manuals include a longitudinal seasonal pattern specifier for major depressive episodes with a consistent temporal relationship with specific times of the year, such as fall and winter, in the past 2 years in individuals with recurrent major depressive disorder (MDD) and bipolar I or II disorder (BD), with the depressive episodes undergoing full remission or, less commonly, a switch to hypomanic or manic symptoms at others times of the year, such as spring or summer.^{3–5} Rosenthal⁵ has argued from many lines of evidence that SAD should be considered an independent clinical entity rather than a seasonal mood specifier. Seasonal affective disorder is characterized by symptoms of depression, such as changes in sleep pattern, fluctuations in weight, decreased energy, and reduced social activities at a particular period of the year, followed by at least partial remission when the season changes.⁶ The most common form is SAD with a winter depression pattern (SAD).² Symptoms of SAD are particularly responsive to bright light treatment.^{7–10} Contrary to common expectations, patients with SAD can manifest severe symptomatology^{11,12} and cognitive deficits¹² similar to nonseasonal depression.¹² Seasonal affective disorder as a clinical diagnosis represents the extreme of a spectrum of seasonality that exists in the population. Many individuals experience seasonal changes in mood and behavior at subclinical levels that can cause significant distress and impairment,^{13,14} while others may experience only very subtle changes.¹⁵

Family studies of SAD prevalence report increased prevalence of SAD in first-degree relatives of patients with SAD.^{16,17} A previous analysis of a large epidemiologic twin study,¹⁸ a subset of which is included in genetic analyses of our data, estimated that genetic factors account for 29% of the overall variance in seasonality in men and women as assessed by the Seasonal Pattern Assessment Questionnaire (SPAQ).¹⁹

Family studies have also shown increased prevalence of nonseasonal depression in the families of individuals with SAD, but the question of whether seasonality and depression are distinct in terms of the genetic risk factors that predispose to them remains unanswered.¹⁵

Table 1. Descriptive Statistics for the Australian and Amish Cohorts

Cohort	Value
Australian sample	
n	3,269
Age, mean $\pm$ SD (range), y	35.7 $\pm$ 13.3 (19–78)
Men, %	34.8
Global seasonality score, mean $\pm$ SD	5.6 $\pm$ 3.9
Old Order Amish	
n	887
Age, mean $\pm$ SD (range), y	55.9 $\pm$ 15.2 (21–99)
Men, %	47.5
Global seasonality score, mean $\pm$ SD	4.5 $\pm$ 3.4

There are differences between MDD with seasonal pattern and BD with seasonal pattern: for instance, the recurrence of the seasonal depression and severity of the course, with higher rates of hospitalization in the bipolar form.²⁰ Importantly, even the unipolar form of SAD has been previously conceptualized clinically as belonging to the bipolar spectrum.²¹

While a number of candidate gene analyses of SAD have been performed,^{22–27} no consistently replicable findings have been gleaned from candidate gene studies of SAD.

Genome-wide association studies (GWASs) have been successful in mapping genetic variants that increase risk of schizophrenia^{28,29} and BD³⁰; however, large sample sizes have been required to detect them. No GWASs have as yet been performed for SAD.

Furthermore, the molecular genetics era has given valuable insights into the etiology of psychiatric disorders that may influence diagnosis in the future. For example, it has recently been demonstrated that much of the genetic risk is shared between psychiatric disorders.³¹

The aim of this study is to perform a GWAS of seasonality in a population of Australian twins and an Old Order Amish population from Lancaster County, Pennsylvania, and to investigate the genetic overlap between seasonality and 3 psychiatric disorders—major depressive disorder (MDD), BD, and schizophrenia.

METHOD

Seasonal Pattern Assessment Questionnaire

A well-studied SAD screening and research questionnaire, the SPAQ³² evaluates SAD by estimating a score of global seasonality calculated by responses on a Likert scale of 0–4 for each of 6 indices of seasonality, the degree to which these changes cause impaired functioning (the problem scale, ranging from 0 to 5), and the season(s) indicated by the participant as representing the time during which they feel “worst.” For cases with incomplete responses on the global seasonality indices, the global seasonality score was estimated from the proportion of responses.

Australian Samples

Participants were drawn from 2 studies conducted between 1988 and 1990 and between 2010 and 2013 at the Queensland Institute of Medical Research. The first study,

- Our study provides evidence for an expected overlap between seasonality of mood and bipolar disorder and a somewhat unexpected genetic overlap between seasonality of mood and schizophrenia. There was no genetic overlap between major depression and seasonality.
- If replicated, our results would imply that seasonal changes in mood and behavior should be evaluated not only in individuals diagnosed with mood disorders but also in patients with schizophrenia. Clinical symptoms that overlap between seasonal affective disorder and schizophrenia include cognitive impairment, social withdrawal, and sleep changes, and both conditions share associations with certain metabolic abnormalities, such as vitamin D deficiency.
- Future research may specifically focus on molecular pathways mediating the overlap between seasonality and schizophrenia and bipolar disorder, potentially leading to theoretical advances and novel therapeutic interventions.

which has been described previously,¹⁸ involved mailing a questionnaire that included the SPAQ to 3,808 twin pairs from the Australian Twin Registry. A total of 2,487 twin pairs and 687 singletons returned questionnaires with responses to the SPAQ. A total of 2,605 individuals provided both genetic and phenotypic information.

The second study, the Midwest Alcohol Research Center study, is a collaborative study between investigators at Queensland Institute of Medical Research and Washington University in St Louis, Missouri, that was initiated to investigate the effects of severe childhood and adult trauma on alcohol consumption and alcohol dependence. The target sample (N=3,607 individuals) had previously participated in a GWAS of these alcohol-related outcomes (Heath et al³³). Participants undertook a phone interview with a trained interviewer with the aim of assessing lifetime history of severe childhood and adult environmental stressors. The SPAQ was included as part of the protocol. At the time of analysis, a total of 686 individuals had completed an interview and provided responses to the SPAQ. After removal of ancestry outliers, a total of 664 individuals provided both phenotype and genotype information.

The overall sample size for inclusion in the GWAS was 3,269. The total sample providing phenotypic information was 6,347. All participants gave informed consent, and the study was approved by the ethics board of Queensland Institute of Medical Research. Descriptive statistics are given in Table 1.

In the interest of reducing the multiple testing statistical burden, it was decided to analyze the global score and not the symptoms individually.

As the Australian twin sample was recruited from different regions of the country, a state of residence fixed effect was included in a linear model along with age and sex, and the residuals of the global seasonality scores after adjusting for these effects were tested for association. Global seasonality scores from all phenotyped individuals were used to extract the residuals.

Genotyping. Full details of the genotyping, imputation, and quality control procedures are given in the supplementary material. After genotyping, quality control, and imputation, a total number of 2,380,486 single nucleotide polymorphisms (SNPs) were included.

Amish

Participants in the study were Amish individuals enrolled in community-based studies^{34–37} conducted between May 2010 and December 2011 at the University of Maryland and University of Maryland's Amish Research Clinic in Lancaster County, Pennsylvania. The participants consented to be contacted for future studies. Seasonal Pattern Assessment Questionnaires were sent by mail to 2,260 such Amish individuals, both male and female and all above the age of 18 years. The letter contained, in addition to the SPAQ, a statement that completion of the SPAQ implied documented informed consent for the study and directions on completion of the questionnaire. This protocol was approved by the institutional review board of the University of Maryland. Initial mailings were done in May 2010. A subsequent mailing in September 2010 was done for those who had not responded to the initial mailing. All responses received before December 31, 2011, numbering a total of 1,306 (response rate 57.8%), comprised the dataset.

Using a convenience subsample of 68 subjects administered the SPAQ twice over a 4-month period, the test-retest reliability of the global seasonality and problem rating score (PRS) in the Amish was adequate using Cronbach α (global seasonality, $\alpha = .87$; PRS, $\alpha = .79$)³⁸ and comparable to that in the general population.³⁹

Genotyping. Genotyping was performed using the Affymetrix GeneChip Human Mapping 500 K or 6.0 Array set (Affymetrix, Santa Clara, California). Genotype calls were made using the Bayesian robust linear model with mahalanobis distance classifier genotype calling algorithm. A total of 364,336 informative autosomal SNPs that passed quality control were included in the analysis. Markov Chain Haplotyping (MaCH)⁴⁰ was used for imputation (release 22, build 36) after applying the following exclusion filters: (1) SNP not found in HapMap; (2) frequency < 0.01 ; (3) Hardy-Weinberg P value $< 1 \times 10^{-6}$; and (4) missingness > 0.05 .

Association Analysis

Association testing in the Australian sample was performed in Merlin,^{41,42} using the dosage scores from the imputation. Merlin accounts for the relationship between individuals in the sample. Four ancestry principal components were included as covariates to correct for population stratification. Prior to this, those individuals with evidence of non-European ancestry were removed from the analysis.

Association analysis in the Amish sample was performed using Mixed Models Analysis for Pedigrees and Populations (MMAP) software developed by J. R. O'Connell (<http://edn.som.umaryland.edu/mmap/index.php>), which accounts for family structure by conditioning the association of genotype with outcome on the relationship matrix (included as a

random effect) and covariates (eg, age and sex). Meta-analysis was performed using the inverse variance weighting method in METAL.⁴³

Statistical Power

We simulated a trait with a heritability of 0.29 using Merlin and estimated the statistical power to detect a variant with a minor allele frequency of 0.2 that explains 1% of the phenotypic variance in global seasonality. The Australian sample had 50.8% power to detect such a variant at the genome-wide significance threshold. The sample had 89% power to detect a variant explaining 1.5% of the phenotypic variance.

Profile Scoring

In the Australian dataset, we tested whether there is a genetic correlation between seasonality, measured by the global seasonality, and mood disorders, utilizing the results from the Psychiatric GWAS Consortium (PGC) GWAS analyses of schizophrenia,²⁸ BD,³⁰ and MDD⁴⁴ to generate genetic profile scores. A single twin from each pair was selected at random for analysis so as to exclude relatives so that the analysis set included 1,999 unrelated individuals. For each individual, 3 separate profile scores were generated based on the results from the PGC GWASs. The profile scoring methodology generates a single genetic "load" score for each individual by weighting each SNP by the log of the odds ratio estimated in the original study. In this way, SNPs with a larger predicted effect on risk to disease are given more weight in making the predictor. Linear regression of the profile score on the trait of interest allows for estimation of how well the profile score predicts the phenotype. This method was first described by Purcell et al⁴⁵ and used to demonstrate that there is overlap in the genetic risk factors for schizophrenia and BD.

We used the summary results from the PGC BD and schizophrenia GWAS analyses separately (downloaded from <http://www.med.unc.edu/pgc>). The schizophrenia study included 9,394 cases and 12,462 controls, and the BP study included 7,481 cases and 9,250 controls. These results had been clumped at $r^2 < 0.25$ to ensure that only the most significant SNP in a given linkage disequilibrium block is included in the analysis and that the same association signal is not included more than once. We also generated profile scores using association results from PGC MDD.⁴⁴ However, since samples from the Queensland Institute of Medical Research contribute to PGC MDD analysis, the PGC MDD data were reanalyzed with the Australian samples excluded, so as to remove any chance of overlap between the discovery and target samples. A total of 7,790 cases and 7,808 controls were included in the revised PGC MDD analysis. The results from this GWAS analysis were used in the MDD profile score analysis.

RESULTS

Genome-Wide Association Study

Results from the most significantly associated SNPs ($P < 10^{-5}$) from the Australian sample are shown in

Supplementary eTable 1. No SNPs reached genome-wide significance in the initial GWAS. All SNPs that had a P value $< 10^{-4}$ and that were independent at $r^2 < 0.5$ or > 50 kilobase (kb) away from each other were tested for replication in the Amish sample. No SNPs passing the significance threshold in the Australian sample were replicated in the Amish with nominal significance ($P < .05$).

A total of 2,354,422 markers that were in common between the 2 studies were included in a meta-analysis of the 2 individual studies. No genes reached the genome-wide significant threshold ($P < 5 \times 10^{-8}$). The most significant SNP in the meta-analysis was rs11825064 ($P = 1.7 \times 10^{-6}$, $\beta = 0.64$, standard error [SE] = 0.13), an intergenic SNP found on chromosome 11. A list of the most significant results is provided in Supplementary eTable 2. A description of the GWAS results and meta-analysis in addition to results from analyses to estimate the heritability explained by all SNPs is reported in Supplementary Methods. Q-Q plots for each of the GWAS analyses are shown in Supplementary eFigures 1–3, and a Manhattan Plot of the results is shown in Supplementary eFigure 4.

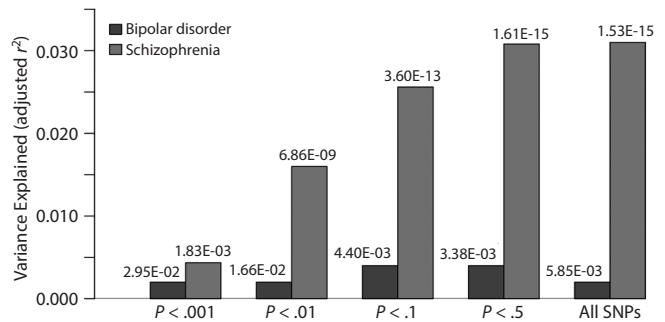
Genetic Overlap Using Profile Scoring

The results for the profile score analysis for all 3 disorders are shown in Supplementary eTable 3. The results for BD and schizophrenia are also shown in Figure 1. There is very strong evidence for genetic overlap between schizophrenia and seasonality ($P < 1 \times 10^{-15}$ for genetic scores based on SNPs associated with schizophrenia at $P < .5$ or all SNPs) and milder evidence for genetic overlap between BD and seasonality ($P = .004$ – $.005$ for genetic scores based on SNPs associated with schizophrenia at $P < .1$, $P < .5$, or all SNPs). The genetic risk score for schizophrenia accounted for up to 3.1% of the phenotypic variation in global seasonality. A nominally significant proportion of the variance in the global seasonality was explained by 427 SNPs, with $P < .001$ from the PGC schizophrenia GWAS ($P = .0018$, variance explained = 0.4%). The null hypothesis that the variance explained by all the schizophrenia SNPs is 0 was rejected ($P = 1.53 \times 10^{-15}$). In contrast, genetic risk score for BD accounted for only 0.4% of the phenotypic variation in global seasonality (see Supplementary eTable 4). No evidence for genetic overlap between MDD and seasonality was observed, as the amount of variance in global seasonality explained by the MDD polygenic scores was not significantly different from 0.

DISCUSSION

We performed a meta-analysis of 2 GWASs of the global seasonality derived from the SPAQ in a sample of twins from Australia and members of the Amish community in Pennsylvania. No genome-wide, significant loci were detected. The results of gene-mapping studies in other affective disorders indicate that a polygenic model, with many common alleles of small effect influencing risk, is

Figure 1. Results From Profile Scoring of Seasonality Using PGC GWAS Results in Australian Sample^a



^aResults from profile scoring analyses using PGC schizophrenia and bipolar GWAS results. The variance explained in seasonality by the profile scores is shown on the y-axis. The significance for each analysis is listed on the top of each bar.

Abbreviations: GWAS = genome-wide association study, PGC = Psychiatric GWAS Consortium, SNP = single nucleotide polymorphism.

likely to explain a significant proportion of the heritability of affected disorders.^{28,45–47} Our results suggest that common variants that are associated with seasonality and that have unusually large effect sizes are unlikely to exist. While the lack of power is a severe limitation of our study, our results will be useful for future meta-analyses of seasonality and SAD, and they provide a list of candidates that can be tested in other cohorts.

The differences between the 2 populations included in this study may have also increased the chances of negative findings. The geographical differences between Australia and the Northeast of the United States are substantial, and there are differences between the amount of sunshine and the day length in different seasons. Even within Australia, there is great variability in day length between cities. We have tried to account for these differences by including state as a covariate in the analysis; however, subtle differences may still exist.

Genetic differences between the populations also exist. The Amish are a genetically isolated population who may harbor unique genetic variants that predispose to seasonality that will not be detected in studies that include other populations. The heritability of seasonality was estimated to be 13.6% in the Old Order Amish,³⁹ a somewhat lower estimate than that in the Australian population-based sample. This indicates that genetic differences between individuals contribute less to the overall variance of seasonality in the Old Order Amish than in Australians. Furthermore, the prevalence of SAD in the Amish is the lowest of all white populations that have been assessed using the SPAQ. This implies that the Amish population is relatively resilient to seasonality. The profile scoring analyses were performed only in the Australian population and may therefore not be generalizable to the Amish. Among Anabaptists there is a low rate of psychosis.^{48,49}

A further limitation of our study is that it did not include actual diagnosis of mental conditions. This would have

permitted the exclusion of those with a psychiatric diagnosis prior to the profile scoring, and therefore the ability to test for genetic overlap between these disorders and seasonality in those without another psychiatric comorbidity. The sample used for profile scoring is a population-based twin sample, so the prevalence of severe psychiatric disorders is likely to be low.

It is noteworthy that genetic profile scores generated from the schizophrenia GWAS explain more variance in seasonality than those derived from the BD and MDD studies. This result is surprising given that seasonal pattern can be added as a modifier of a unipolar and bipolar depression diagnosis in *DSM-5*. In contrast, the seasonal pattern is not clinically or epidemiologically considered in association with schizophrenia. The finding that the schizophrenia profile scores explain a larger proportion of variance in the global seasonality may be due to differences between the GWASs from which the polygenic scores were generated. For instance, a larger sample size or better accuracy of diagnosis or lack of some other confounding source could lead to more accurate estimates of the true SNP effects in the schizophrenia GWAS compared to those of BD or MDD. Summing more accurate SNP effect estimates of over thousands of SNPs could lead to substantially more accurate polygenic scores.

In contrast to BD, MDD showed no evidence for genetic overlap with global seasonality in our sample. Conceptually, SAD has been considered as a condition on the spectrum of bipolarity,²¹ even if major depression episodes alternate with periods of remission. To date, GWASs of MDD have failed to uncover any replicable common variants. One possible reason for this is that the estimates of the SNP effects from the MDD GWAS may be less accurate than those from the schizophrenia and BD GWASs. It has been argued that a GWAS of MDD has less power than a GWAS of schizophrenia or BD of the same sample size, simply reflecting that MDD is a more prevalent disorder, is potentially more heterogeneous, and has a lower proportion of risk explained by genetic factors.⁵⁰

Although seasonality of mood and general health has been observed since ancient times, seasonality of psychiatric symptoms (predominantly affective, but also psychotic and catatonic features as well as cognitive deficits) was clinically described in the early 19th century by Pinel and his student Esquirol, with additional early contributions by Griesinger, Kraepelin, and Kraines (reviewed by Wehr and Rosenthal¹). To our knowledge, our current report provides the first direct evidence at the molecular level for the overlap between seasonality and BD that has been previously suggested clinically.²¹

While the association between seasonality of mood and schizophrenia was unexpected, a hint of a relationship was reported in small samples at high latitudes.⁵¹ Specifically, Doorack et al⁵¹ examined SAD symptoms in patients with schizophrenia in Alaska and found that 36% had co-occurring SAD. Moreover, previous studies have suggested an association between onset of the first

episode of schizophrenia and season (eg, Strous et al⁵²). Additionally, studies conducted in the northern hemisphere found a summer peak of schizophrenia admissions in hospitals.^{53–55} In England and Wales, for example, Hare and Walter⁵⁴ found a summer peak of schizophrenia admissions. Clarke et al⁵⁶ found a seasonal association in first-episode schizophrenia admissions in Ireland, although this association varied on an annual basis. In a subsequent study, Clarke et al⁵⁷ found that the seasonal association held only for first admissions and not for subsequent admissions. Shiloh et al⁵⁸ found a summer peak for schizophrenia admissions in Israel. In China, Tian et al⁵⁹ found a spring peak (March) in schizophrenia admissions, although they noted that they did not distinguish between first episodes and readmission, which, according to Clarke et al,⁵⁷ may be necessary to isolate seasonality effects. However, there are studies that found no seasonal variation in schizophrenia or that found patterns contrary to those previously reported. For example, Aviv et al⁶⁰ as well as Amr and Volpe⁶¹ did not find any seasonal effect on admission rates. In the southern hemisphere, Davies et al⁶² found a peak in first-episode schizophrenia admission in the winter (August), similar to Owens and McGorry.⁶³ In contrast, Daniels et al⁶⁴ did not find a seasonal association in admissions with diagnoses of schizophrenia and bipolar disorder in Tasmania.

There are certain clinical symptoms that overlap between SAD and schizophrenia, including cognitive deficits, social withdrawal, and changes in sleeping pattern. Furthermore, the incidence of schizophrenia has been found to be higher at higher latitudes,⁶⁵ a pattern that is shared with SAD, and it has been shown that there is a season of birth effect for schizophrenia,⁶⁶ and likewise for SAD,⁶⁷ which has led to the hypothesis that vitamin D plays a crucial role in the etiology of schizophrenia.⁶⁸ Vitamin D improves mood in healthy individuals⁵¹ during winter and reduces depression scores in patients with SAD.⁵² The results of this study suggest that, in addition to environmental modifiable risk factors such as ultraviolet B radiation, photoperiod, visible light intensity, vitamin D supplementation and levels, skin exposure (eg, sunscreen, reduced outdoor activities), and weight increase (resulting in lowering of vitamin D levels^{69,70}), SAD and schizophrenia share previously unacknowledged genetic risk factors that deserve studies in their own right. However, our study provides evidence only for an overlap in genetic risk factors between schizophrenia and seasonality. Further evidence for an association between seasonality and schizophrenia is needed at the genetic, environmental, and clinical levels before seasonality can be considered as a component of schizophrenia.

In conclusion, we provide direct evidence for an expected genetic overlap in risk between bipolar disorder (but not major depressive disorder) and seasonality, and a somewhat less expected overlap between schizophrenia and seasonality. Further investigation of the links between bipolar disorder, schizophrenia, and SAD at both the clinical and molecular levels is warranted and may lead, in the long run, to studies that uncover novel therapeutic targets.

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genotype data can be found at www.med.unc.edu/pgc. Psychiatric GWAS Consortium (PGC) contributors and affiliations are listed in Supplementary eTable 4.

Supplementary material: Available at PSYCHIATRIST.COM.

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Supplementary Material

Article Title: Seasonality Shows Evidence for Polygenic Architecture and Genetic Correlation With Schizophrenia and Bipolar Disorder

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Supplementary Material

Supplementary eTable 4. List of contributors to PGC-MDD GWAS study

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Supplementary Methods

QIMR Genotyping and Imputation

Genotype information was collected as part of a number of genotyping projects undertaken at QIMR. DNA samples were collected in accordance with standard protocols and submitted to different genotype centres using different SNP platforms (Illumina 317K, IlluminaHumanCNV370- Quadv3, and Illumina Human 610-Quad). SNPs were called using the Illumina BeadStudio software. A standard quality control (QC) procedure was applied to each project individually, prior to imputation. A detailed description of the QC steps and procedure for detection of ancestry outliers is given elsewhere (Medland et al., 2009).

A set of 274,604 SNPs that were common to all of the genotyping chips was used for imputation, which was performed using the program MACH.(Li et al., 2009) The imputed SNPs were screened further for Mendelian errors, minor allele frequency and missingness. Only SNPs with an imputation quality score (R^2) greater than 0.3 were retained,

Results of GWAS meta-analysis

There was little evidence of population stratification in any of the analyses. The lambda inflation factor for the GWAS in the Australian population was 1.00, and the same value was found for the meta-analysis of the Australian and Amish results. Lambda was equal to 1.03 in the Amish analysis (Supplementary Figure 1-3).

The most significant SNP in the meta-analysis was rs11825064 ($p = 1.7 \times 10^{-6}$, $\beta = 0.64$, S.E = 0.13), an intergenic SNP found on chromosome 11. This region has not been implicated in mood disorders previously. Other SNPs that show the greatest evidence of

association with the GSS include rs13257657, a SNP found in a region encoding a microRNA (mir-1204). The microRNA.org software predicts that mir-1204 may potentially target a wide range of genes, (Betel et al., 2008), with the strongest evidence for the *MYLK* gene as a target.

Another SNP - rs1808478 – also shows evidence of association ($p = 4.7 \times 10^{-4}$, $\beta = -0.59$, S.E. = 0.13. This SNP is found in the *POLR2M* gene, which encodes a subunit of the RNA polymerase II complex. This gene is also a homolog of the ionotropic glutamate receptor N-methyl D-aspartate, indicating that it may play a role in signalling the brain.

Estimating Phenotypic Variance Explained by all SNPs

A whole-genome method that estimates the genetic variance explained by all SNPs on the chips (Yang et al.), implemented in the freely available GCTA software (Yang et al., 2011) was utilised. The Amish sample was not included in this analysis. Initially, one individual per family was selected at random from the Australian sample ($n = 1,999$) individuals. Only genotyped SNPs were utilised. A set of quality control steps were implemented to remove SNPs with high rates of missingness or Hardy-Weinberg Equilibrium p -value < 0.001 . After quality control, a total of 217,053 SNPs remained. Individuals with more than 1% missing SNPs ($n = 44$) were removed. To ensure that there was no cryptic relatedness in the sample that could bias the results, individuals with a pairwise genetic relatedness coefficient as estimated from the SNPs of more than 0.025 were removed. After QC steps, a total of 1,864 individuals remained. The estimated proportion of the variance explained was 8.6% with a standard error of 20.0%. Including the genetic component estimated from the SNPs did not significantly improve the fit of the model, when compared to including only environmental effects ($p = 0.30$). The large

standard error of the estimate indicates that larger sample sizes are needed in order to have adequate power to reject the null hypothesis that common variants on the SNP chips explain a proportion of the heritability of seasonality.

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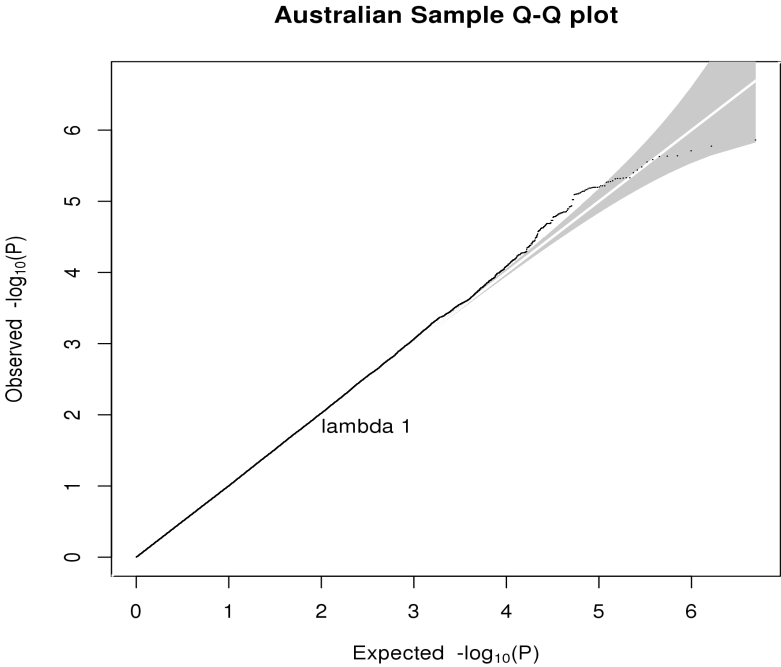
Supplementary eTable 1. The most significant SNPs in the Australian sample and their corresponding results in the Amish sample.

SNP	P	CHR	coordinate	Tested Allele	Freq	Beta	S.E.	closest gene	distance to gene	MAF in Amish	Beta Amish	S.E. Amish	P Amish
rs12411769	1.38E-06	10	124969001	C	0.04	1.87	0.39	LOC100131719	-18669	0.02	0.17	1.65	0.91
rs9555488	1.67E-06	13	109254277	T	0.29	-0.54	0.11	MYO16	0	0.13	-0.03	0.29	0.91
rs1813849	1.95E-06	2	200142847	T	0.12	0.81	0.17	SATB2	0	0.07	-0.69	0.43	0.11
rs13257657	4.72E-06	8	128986335	C	0.13	0.84	0.18	MIR1204;PVT1	0	0.04	-0.79	0.57	0.16
rs31019	5.14E-06	5	96422368	T	0.45	0.50	0.11	LIX1	-2411	0.46	-0.09	0.20	0.66
rs7870657	5.37E-06	9	71724126	G	0.38	0.49	0.11	FXN	9032	0.49	-0.19	0.19	0.31
rs2498436	6.50E-06	9	71790905	G	0.29	0.52	0.12	TJP2	0	0.32	0.03	0.20	0.89
rs1808478	7.18E-06	15	58009759	T	0.18	-0.60	0.13	GCOM2	0	0.10	-0.48	0.58	0.41
rs3936510	7.22E-06	5	55860866	T	0.19	0.59	0.13	OTTHUMG000000059482	0	0.16	-0.02	0.27	0.93
rs379533	7.83E-06	1	30483596	T	0.46	-0.47	0.11	OTTHUMG000000003683	3203	0.35	0.34	0.20	0.08

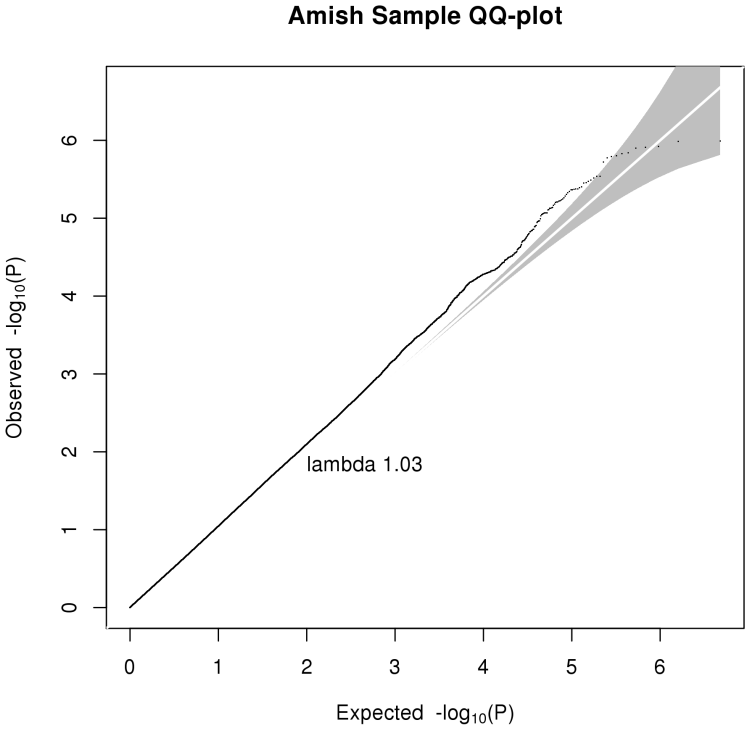
Supplementary eTable 2. Independent SNPs with $p < 10^{-5}$ from meta-analysis of cohorts

SNP	CH R	coordinate	Meta_ Pval	Meta_ Beta	META_ SE	closest gene	distance to gene	Tested Allele	Non-tested Allele	Directi on	Freq_A mish	Freq_QI MR	Pval_Amish	Pval_QIMR
rs11825064	11	134477854	1.70E-06	0.64	0.13	OTTHUMG00000165125	34904	c	g	++	0.19	0.15	2.30E-04	9.28E-04
rs13257657	8	128986335	1.73E-06	-0.83	0.17	MIR1204;PVT1	0	t	c	--	0.04	0.13	1.61E-01	4.72E-06
rs12411769	10	124969001	2.30E-06	-1.78	0.38	OTTHUMG00000019196	-37606	a	c	--	0.98	0.96	9.18E-01	1.38E-06
rs1808478	15	58009759	4.68E-06	-0.59	0.13	GCOM2;POLR2M	0	t	c	--	0.90	0.82	4.09E-01	7.18E-06
rs9559377	13	109264870	5.01E-06	0.49	0.11	MYO16	0	a	t	++	0.86	0.73	6.88E-01	2.80E-06
rs196889	8	24734677	6.40E-06	0.68	0.15	OTTHUMG00000163790	0	t	c	++	0.87	0.91	1.68E-03	9.44E-04
rs1131351	2	20402380	7.51E-06	0.51	0.11	SDC1	0	c	g	++	0.66	0.59	1.31E-02	1.54E-04
rs7870657	9	71724126	8.56E-06	-0.41	0.09	FXN	9032	a	g	--	0.49	0.62	3.09E-01	5.37E-06

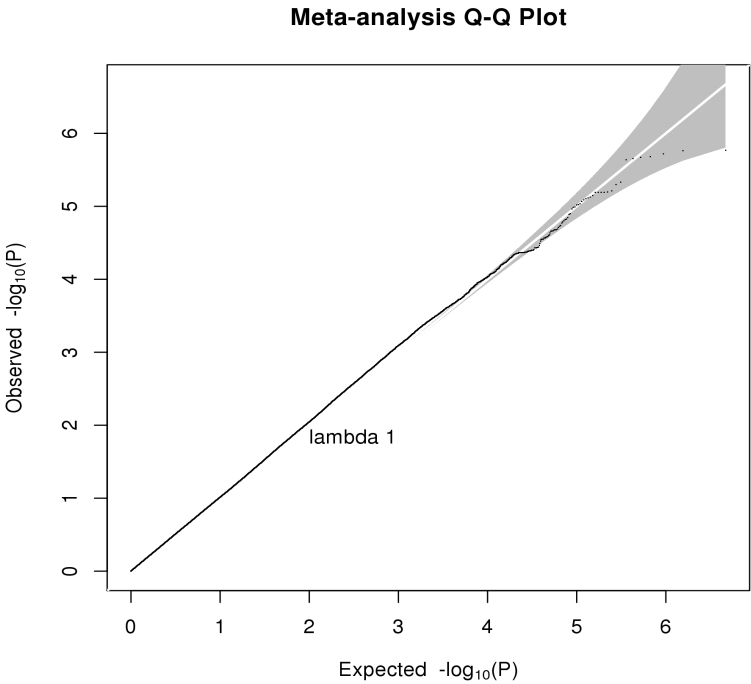
Supplementary eFigure 1. Q-Q plot from Australian GWAS analysis



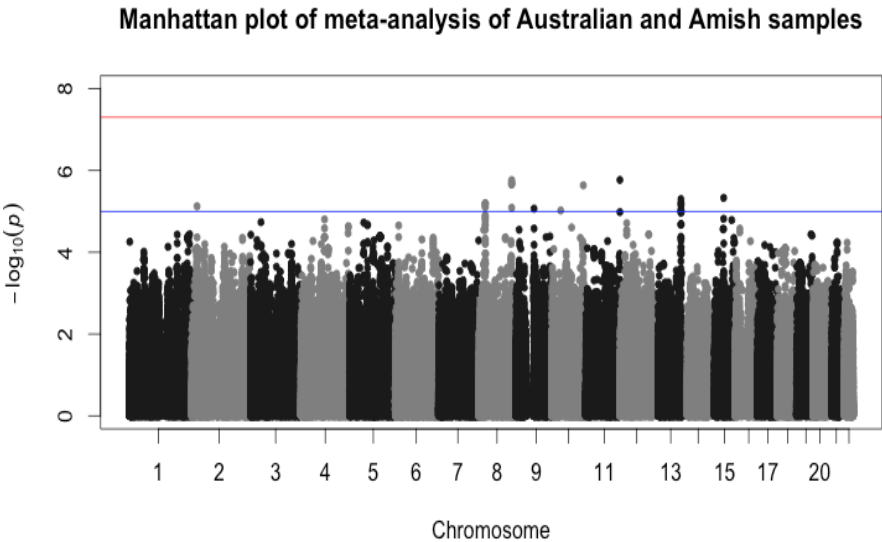
Supplementary eFigure 2. Q-Q plot from GWAS analysis in Amish sample



Supplementary eFigure 3. Q-Q plot for meta-analysis



Supplementary eFigure 4. Manhattan Plot for meta-analysis



Supplementary eTable 3. Results of profile scoring analyses using results from Psychiatric Genetics Consortium GWAS of schizophrenia, bipolar disorder and Major Depressive Disorder

BIP SNP significance cutoff	No. of Alleles	Adjusted R2	p-value for model
p < 0.001	737	0.002	2.95E-02
p < 0.01	4060	0.002	1.66E-02
p < 0.1	24096	0.004	4.40E-03
p < 0.5	77395	0.004	3.38E-03
All SNPs	108832	0.003	5.85E-03
MDD SNP significance cutoff	No. of Alleles	Adjusted R2	p-value for model
0.001	310	0	0.52
0.01	2743	0	0.25
0.1	20586	0	0.10
0.5	83047	0	0.32
All SNPs	154538	0	0.29
SCZ SNP significance cutoff	No. of Alleles	Adjusted R2	p-value for model
0.001	427	0.004352352	1.83E-03
0.01	3348	0.016183205	6.86E-09
0.1	23610	0.025636295	3.60E-13
0.5	83903	0.0308183	1.61E-15
All SNPs	123040	0.030867498	1.53E-15